

1 A factor analysis of Functional Independence and Functional Assessment Measure scores
2 among focal and diffuse brain injury patients: The importance of bi-factor models.

3

4 Sarah Gunn (MSc), Gerald H. Burgess (PsyD), and John Maltby (DPhil).

5

6 Affiliation: Neuroscience, Psychology and Behaviour, College of Life Sciences, University
7 of Leicester, Leicester.

8

9 Correspondence concerning this article should be addressed to Sarah Gunn, Neuroscience,
10 Psychology and Behaviour, College of Life Sciences, University of Leicester, Leicester, LE1
11 9HN, United Kingdom. Tel: +44 (0)116 252 2482. Email address: scg30@le.ac.uk.

12

13 This research did not receive any specific grant from funding agencies in the public,
14 commercial, or not-for-profit sectors.

15

16

17

18

19

20

21

22

23

24

25

26 **A factor analysis of FIM+FAM scores among focal and diffuse brain injury patients:**

27 **The importance of bi-factor models.**

28

29

30

31 **Structured Abstract**

32

33

34 **Objective:** To explore the factor structure of the UK Functional Independence Measure and
35 Functional Assessment Measure (FIM+FAM) among focal and diffuse acquired brain injury
36 patients.

37 **Design:** Criterion standard.

38 **Setting:** An NHS acute acquired brain injury inpatient rehabilitation hospital.

39 **Participants:** Referred sample of 447 adults (835 cases after exclusions) admitted for
40 inpatient treatment following an acquired brain injury significant enough to justify intensive
41 inpatient neurorehabilitation.

42 **Intervention:** Not applicable.

43 **Outcome measure:** Functional Independence Measure and Functional Assessment
44 Measure.

45 **Results:** Exploratory Factor Analysis suggested a two-factor structure to FIM+FAM scores,
46 among both focal-proximate and diffuse-proximate acquired brain injury aetiologies.

47 Confirmatory Factor Analysis suggested a three-factor bi-factor structure presented the best
48 fit of the FIM+FAM score data across both aetiologies. However, across both analyses, a
49 convergence was found towards a general factor, demonstrated by high correlations between
50 factors in the Exploratory Factor Analysis, and by a general factor explaining the majority of
51 the variance in scores on Confirmatory Factor Analysis.

52 **Conclusion:** Our findings suggested that although factors describing specific functional
53 domains can be derived from FIM+FAM item scores, there is a convergence towards a single
54 factor describing overall functioning. This single factor informs the specific group factors
55 (e.g. motor, psychosocial and communication function) following brain injury. Further

56 research into the comparative value of the general and group factors as evaluative/prognostic
57 measures is indicated.

58

59 **Keywords:** brain injuries; rehabilitation; treatment outcome; factor analysis

60

61 **Abbreviations:**

62 Acquired Brain Injury (ABI)

63 Comparative Fit Index (CFI)

64 Confirmatory Factor Analysis (CFA)

65 Exploratory Factor Analysis (EFA)

66 Functional Independence Measure and Functional Assessment Measure (FIM+FAM).

67 Keiser–Meyer–Olkin (KMO)

68 Non-normed fit index (NNFI)

69 Relative chi-square degrees of freedom (CMIN/DF)

70 Root mean square of approximation (RMSEA)

71 Standardised root mean square residual (SRMR)

72

73

74

75

76

77 Patients with moderate to severe acquired brain injury (ABI) may experience long-lasting or
78 permanent difficulties with mobility, activities of daily living, cognition and social
79 reintegration¹. Accurate functional assessments enable interdisciplinary teams to set
80 meaningful rehabilitation goals, make better predictions about prognosis, and identify
81 appropriate discharge placements earlier in rehabilitation^{2,3}.

82

83 The UK Functional Independence Measure and Functional Assessment Measure
84 (FIM+FAM⁴) is used in complex ABI rehabilitation services UK-wide and internationally. It
85 evaluates functional impairment and assistance needs across physical, communication and
86 psychosocial domains, using input from the interdisciplinary therapeutic team. The
87 FIM+FAM is a reliable, valid scale with high internal consistency, excellent test-retest
88 reliability and very good inter-rater reliability, and is one of the most widely-used outcome
89 measures in ABI rehabilitation^{2,4-7}. However, conflicting arguments have been advanced
90 regarding the structure of the FIM+FAM, affecting its interpretation and prognostic utility.

91

92 Initial key research supported a two-factor FIM+FAM structure, comprising a motor
93 and cognitive subscale as per the manual^{8,9}. However, further work suggested a greater
94 number of factors may better explain the variance in scores. For example, among a general
95 neurorehabilitation sample, a four-factor FIM+FAM structure was identified incorporating a
96 motor factor (comprising 15 of the 16 original motor items), subdivision of the cognitive
97 scale into psychosocial (9 items, e.g. social interaction and emotional status) and
98 communication elements (5 items, e.g. comprehension and expression), and the final factor
99 comprised 6 activities of daily living items, plus the community mobility item formerly
100 viewed as part of the motor subscale¹⁰. Similarly, among stroke patients a three-factor
101 structure was identified comprising 15 of the 16 original motor items (excluding swallowing),

102 and the same division into psychosocial and communication factors; this was superior to the
103 two-factor model comprising motor items and the broader cognitive factor¹¹.

104

105 These analyses indicate the validity of a multifactorial interpretation of the
106 FIM+FAM, with factor structures demonstrating specific and independent dimensions of
107 function identifiable on assessment following ABI⁷. However, both of the aforementioned
108 studies also reported salient loading of the FIM+FAM items onto a single component; this
109 suggests a potential additional use of the scale as a measure of general functioning, aside
110 from the more faceted multifactorial solutions⁷. This implies the possible validity of a bi-
111 factor model solution, which integrates single and multiple factor solutions. This enables
112 retention of a single common construct (e.g. general functional independence), while also
113 acknowledging multidimensionality (e.g. motor, psychosocial and communication group
114 factors)¹². However, previous examinations of potential factor structures of the FIM+FAM
115 have not considered a potential bi-factor model^{7,11}.

116

117 The division of the FIM+FAM into multiple factors of function is also worth
118 exploring in the context of varied injury aetiologies, which can present differently clinically
119 and thus differentially influence prognosis, care management and rehabilitative input. For
120 example, previous research FIM+FAM research found different functional outcomes between
121 right- and left-sided stroke patients¹¹. Similar differences may exist between the effects of
122 focal and diffuse brain injury. Focal injury is generally limited to a smaller, more defined
123 area and is typically associated with greater physical impairment and fewer cognitive effects,
124 whereas patients with diffuse injuries often retain more physical function, but with greater
125 cognitive impairment (particularly regarding communication and psychosocial functions)^{13,14}.
126 However, many ABI comprise elements of both (e.g. diffuse axonal injury resulting from

152 unwell to engage in rehabilitation). Patients were also excluded who had a non-ABI diagnosis
153 (e.g. pain syndromes, psychiatric disorders or severe physical trauma not involving the brain),
154 progressive conditions (e.g. multiple sclerosis), or other rare conditions where inclusion
155 would have compromised patient anonymity.

156

157 The overall sample composition in terms of injury type was 40% trauma, 37.4%
158 stroke, 14.1% anoxia, 5.1% inflammatory condition and 3.4% tumour (or injury by tumour
159 removal). In terms of localisation, for 6.2% of patients, injury location had not been
160 documented and these cases were therefore not included in analysis. Of the remainder, 49.4%
161 of patients had sustained global, diffuse or bilateral-hemisphere injury and 50.6% had
162 sustained a localised injury to the frontal region, the brainstem/cerebellum, or the left/right
163 hemisphere.

164

165

166 **Measure**

167

168

169 All inpatients were assessed for cognitive and functional impairments using the 36-item
170 English-language FIM+FAM on admission and discharge⁹, comprising assessments of self-
171 care, mobility, communication, cognition, mood and social behaviour. Each patient's
172 admission FIM+FAM was completed by an interdisciplinary team (allocated consultant,
173 clinical psychologist, physiotherapist, occupational therapist, speech and language therapist
174 and dietitian) meeting two weeks post-admission, describing the patient's
175 impairment/function on arrival. The follow-up FIM+FAM was completed at the first weekly
176 meeting post-discharge. Most FIM+FAM items are scored between 1-7 (except item 14.2,

177 wheelchair locomotion, which is scored 0-6; 0 indicates that the patient never requires a
178 wheelchair), with 7 indicating total independence and decreasing scores indicating greater
179 impairment. Demographic and aetiological data were collected from medical records.

180

181 The “Extended Activities of Daily Living” section (items 31-36)⁶ was excluded from
182 this analysis because it had been inconsistently completed over the years. 80.6% (5734/7111)
183 of ratings were scored at 1, which is both the lowest score possible and the default rating
184 when assessment had not yet been completed, with no means to differentiate between which
185 scores were accurate and which were placeholders.

186

187

188 **Data Preparation, Missing Values and Analysis**

189

190

191 Consistent with previous research⁷, we included FIM+FAM scores acquired upon
192 participants’ admission and discharge to maximise the range of scores sampled across the
193 population. This increased the sample size to $n=894$. However, some cases were removed.
194 Seven cases were removed as the UK FIM+FAM allows a ‘0’ score for wheelchair
195 locomotion if a wheelchair is never used, making this data incongruous with the rest of the
196 scoring. Fifty-two cases were removed due to missing injury localisation data. 835 cases were
197 taken forward, from which 420 cases had a focal-proximate brain injury (stroke or trauma)
198 and 415 cases had a diffuse-proximate brain injury (anoxia or inflammatory condition).

199

200

201 **Ethics**

202

203

204 The South Warwickshire NHS Clinical Audit and Effectiveness Department (registration:
205 1400) and the University of Leicester (reference 9256) provided approval. Full ethical board
206 review was not required, since no additional data were collected.

207

208

209

Results

210

211

212 Exploratory Factor Analysis (EFA) is used to discover the underlying structure of items
213 within a data set. Confirmatory Factor Analysis (CFA) is used to test whether proposed
214 structures to a set of items provide a good explanation of the data. No previous research has
215 reported on the factor structure for FIM+FAM item scores among focal-proximate and
216 diffuse-proximate brain-injured individuals, and the clinical presentation of symptoms is
217 complex. Therefore, we subjected the data to both EFA (to discover the underlying structure
218 of the items) and CFA (to test possible structures to the set of items). Accordingly, after
219 removing missing cases, we divided both the focal-proximate and diffuse-proximate
220 participant data into two samples (focal-proximate, $n=210/210$; diffuse-proximate,
221 $n=207/208$) using SPSS for Windows™ randomly-generated numbers to place participants in
222 a random order and assign them to the EFA or CFA sample.

223

224

225 Exploratory Factor Analysis

226

227
228 Across the focal-proximate and diffuse-proximate data used for EFA, the ranges for skewness
229 and kurtosis statistics for the 30 FIM+FAM items fell between -1.25 to 1.46 (skewness) and -
230 1.91 to 0.36 (kurtosis). These statistics fall within criteria of values within +/-2 representing
231 "acceptable" symmetry¹⁶⁻¹⁸. Consequently, an EFA with maximum likelihood extraction with
232 promax rotation was conducted with the assumption of normality of the data.

233
234 Both participant sample size (focal-proximate, $n=210$; diffuse-proximate, $n=208$) to
235 variables (30) ratios exceeded the minimum recommended ratio for EFA of 5 to 1, with a
236 minimum participant sample of 150¹⁹. Bartlett's test confirmed that an EFA was appropriate
237 for the focal-proximate sample ($\chi^2[435]=8290.70, p<.001$) and diffuse-proximate sample
238 ($\chi^2[435]=11068.50, p<.001$). A Keiser–Meyer–Olkin (KMO) test indicated there was a
239 sufficient participant:item ratio for both the focal-proximate sample (7:1, KMO=.96) and the
240 diffuse-proximate sample (6.93:1, KMO=.96).

241
242 Parallel analysis was used to determine the number of factors to extract, based on
243 findings suggesting that this method (in which eigenvalues are compared to those expected
244 from purely random data) is the most appropriate and accurate^{20,21}. For the focal-proximate
245 sample, the third eigenvalue (17.71, 3.91, 1.50) failed to exceed the third mean eigenvalue
246 (1.79, 1.67, 1.59) calculated from 1,000 generated datasets with 210 cases and 30 variables,
247 suggesting a two-factor solution. For the diffuse-proximate sample, the third eigenvalue
248 (21.31, 2.81, 1.03) also failed to exceed the third mean eigenvalue (1.79, 1.67, 1.59)
249 calculated from 1,000 generated datasets with 208 cases and 30 variables, again suggesting a
250 two-factor solution.

251

252 A two-factor solution was therefore sought for both samples using a promax rotation,
253 as the factors were anticipated to be correlated, with delta set to 0 (Table 1). Meaningful
254 loadings were assessed using the criteria of 0.32 (*Poor*), 0.45 (*Fair*), 0.55 (*Good*), 0.63 (*Very*
255 *good*) or 0.71 (*Excellent*)²², with a minimum of three items loading significantly on any
256 element to confirm it as an independent factor^{23,24}. Both solutions are best described by the
257 original two-factor model, comprising motor and cognitive factors. However, for the focal-
258 proximate sample, one supposedly motor item ('swallowing') loaded on the cognitive factor
259 rather than motor, despite being theoretically connected to the latter. Furthermore, among
260 both samples, both factors were highly correlated; focal-proximate, $r=.62$, and diffuse-
261 proximate, $r=.73$.

262 - Insert Table 1 here -

263

264

265 **Confirmatory Factor Analysis**

266

267 To explore the structural validity of the FIM+FAM, a series of comparisons using CFA was
268 performed using AMOS 24 software with the second sets of randomly-assigned samples
269 (focal-proximate, $n=210$; diffuse-proximate, $n=207$). Though evaluating acceptability of
270 model fit against key criteria is a major focus of CFA, it is additionally useful for
271 demonstrating the incremental value of proposed models²⁵. This is important for the current
272 consideration, which seeks not to exclude items, but to understand how best to conceptualise
273 the relationships between the variables. Six possible models were tested for goodness-of-fit.
274 The first was the proposed two-factor structure comprising motor and cognitive components⁸,
275 which incorporates our findings from EFA. The second structure was a three-factor model¹¹,
276 comprising motor (15 items, with the 'swallowing' item excluded), psychosocial (9 items)

277 and communication factors (5 items). The third structure was the four-factor model⁷,
278 comprising physical (15 items, without the community mobility item), psychosocial (9
279 items), communication (5 items) and activities of daily living (6 items, plus community
280 mobility) factors. The fourth proposed structure was a unidimensional model representing an
281 underlying latent factor structure of general functioning. The fifth, sixth and seventh
282 structures were bi-factor versions of the two-, three- and four-factor models.

283

284 To examine the goodness-of-fit of the data against key criteria, we used the following
285 recommended statistics^{26,27}: the chi-square (X^2), the comparative fit index (CFI), the non-
286 normed fit index (NNFI) and the root mean square error of approximation (RMSEA).
287 Additionally, we report the relative chi-square degrees of freedom (CMIN/DF). We used the
288 following criteria to assess whether the model fit was adequate (noting the chi-square test was
289 likely to be significant due to the large sample size²⁸); (i) the relative chi-square (CMIN/DF)
290 should be less than 3 to be acceptable, and less than 2 to be 'good', (ii) the CFI and NNFI
291 should exceed .90 to be acceptable and exceed .95 to be 'good' and (iii) the RMSEA should
292 not exceed .08, and should be below .06 to be a 'good' fit^{27,29}. In terms of improved fit for
293 models, we assessed improved goodness-of-fit by changes in CFI (Δ CFI) being $>.01$ ³⁰.

294

295 Table 2 shows goodness-of-fit statistics for the seven models. Among both the focal-
296 proximate and diffuse-proximate samples, nearly all the goodness-of-fit statistics did not
297 meet all the aforementioned criteria for acceptability (noting that the SRMR is unobtainable
298 for the four-factor and corresponding bi-factor model⁷, due to one factor comprising one
299 item). There was one exception; the three-factor¹¹ bi-factor model analysis met the
300 acceptability criteria for goodness-of-fit statistics, where the CFI statistic exceeded .90. In
301 terms of improvement of fit for CFI statistics obtained compared to other models²⁵, as

302 indicated by changes in ΔCFI being $>.01$ ³⁰, the three-factor bi-factor model proved the better
303 fit among both the focal-proximate and diffuse-proximate samples than the six other
304 proposed models. The variance accounted for by the general factor in the three-factor bi-
305 factor model was 74.7% and 80.4% for the focal-proximate and diffuse-proximate samples
306 respectively. The variance accounted for by the motor, psychosocial and communication
307 factors were, respectively, 6.8%, 12.0% and 6.6% for the focal-proximate sample, and 3.2%,
308 10.8% and 5.6% for the diffuse-proximate sample.

309 - Insert Table 2 here -

310

311

312 Discussion

313

314

315 The current study suggested potential validity of a more general interpretation of FIM+FAM
316 scores. We examined FIM+FAM factor structures among patients with focal-proximate and
317 diffuse-proximate ABI, comparing single versus multifactorial solutions, with the assumption
318 that the former could offer greater clinical utility in some situations. The EFA suggested a
319 two-factor solution consistent with the original scoring of the scale⁹, while the CFA
320 suggested a three-factor¹¹ bi-factor solution presented the best fit. However, our EFA and
321 CFA suggest a weighting towards a single general factor. For the EFA, in both focal-
322 proximate and diffuse-proximate brain injury patients, the loadings for some items were
323 above 1, suggesting a high degree of multicollinearity between the items³¹, and the
324 correlations between the factors were large³² ($r >.62$). For the CFA, the general factor also
325 accounted for a high degree of variance ($>74.7\%$ across both samples), suggesting that the
326 variance for the items was explained by the general factor. This may have implications for

327 conceptualisation of FIM+FAM scores; our findings suggest a higher-order structure to these
328 items, with the general factor underpinning overall functioning and informing group factors.
329 This finding is consistent across the two major aetiological groupings of focal-proximate and
330 diffuse-proximate brain injury.

331

332 This evidence for a single general factor contrasts with recent findings suggesting that
333 the FIM+FAM comprises multiple factors^{8,11,33}. These differences might relate to various
334 issues. Firstly, the current study omitted the Extended Activities of Daily Living component
335 due to poor data quality, which might explain differences in derived structure from previous
336 research. Secondly, the timeframe for scoring differed across studies; for example, a previous
337 study's⁸ scores were obtained within 48 hours of admission, while the current study's scores
338 were generated within 10 working days consistent with manualised administration⁹. This
339 provided more time for teams to assess admissions, which may have generated differences in
340 scoring. Finally, cohorts differ between studies, which may have produced differential
341 outcomes. Previous research has assessed factor structure specifically with traumatic brain
342 injury⁸ and stroke patients¹¹, while this study utilised data from an inclusive sample of
343 patients with trauma, stroke and other acquired aetiologies; this is representative of typical
344 cohorts assessed using the FIM+FAM.

345

346 The finding of a general factor informing group factors in a bi-factor model¹² presents
347 a different theoretical proposition to the currently-dominant view that the FIM+FAM
348 generates specific and independent factors (e.g. motor, psychosocial and communication)
349 describing function post-injury. Clearly, the ability to assess specific domains in brain injury
350 outcomes is crucial to evaluate differential progress, to generate appropriate rehabilitative
351 goals, and to make realistic prognostic predictions³⁴. However, availability of a general factor

352 of functioning which provides an equivalent (or better) summary of overall impairment may
353 also be useful, as simpler models tend to be more helpful and pragmatically applicable in
354 clinical settings³⁵. An immediate target for future study would be to explore the
355 evaluative/prognostic utility of the general versus specific conceptualisations in bi-factor
356 FIM+FAM models – particularly given the variance in prior research in terms of time-frames,
357 measures used, brain injury aetiologies and inclusion (or not) of the Extended Activities of
358 Daily Living component. The general factor providing a useful model for assessment and
359 prognosis would be theoretically and clinically important in rehabilitation.

360

361

362 **Limitations**

363

364

365 Distinguishing between focal-proximate and diffuse-proximate injury is important; many
366 ABIs combine elements of both, and while we attempted classification via broad
367 categorisations, the clinical delineation is not always clear. This may have affected the
368 derived factor structures. In this retrospective analysis, detailed data was unavailable to
369 classify injuries more accurately as focal/diffuse; however, future studies should consider
370 acquiring/using this information.

371 It is also important to consider that including only patients with very complex injuries
372 both limits generalisability to those with less complex injuries, and may have masked
373 differences in functional ability which could potentially be more evident in those with less
374 generalised/complex impairment.

375 The lack of good-quality Extended Activities of Daily Living data limits
376 comparability with some past research; future studies should discriminate between minimum

377 scores denoting actual minimum function, versus no assessment. Additionally, pooling
378 admission/discharge data for factor analysis risks high intercorrelation between scores.
379 Finally, generalisability is limited when multivariate models are developed/tested at one
380 rehabilitation unit³⁷; confirmatory studies from additional sites are required.

381

382

383

Conclusions

384

385

386 This study reports the first factor analysis of FIM+FAM scores to draw a distinction between
387 focal-proximate and diffuse-proximate brain injury, and to test bi-factor models. Our findings
388 suggested that although independent factors can be derived from FIM+FAM item scores,
389 there is a convergence towards a factor describing overall functioning, which additionally
390 informs specific group factors following brain injury. This may with further study prove to be
391 of significant clinical utility.

392

393

394

References

- 395
- 396
- 397
- 398 1. Shames J, Treger I, Ring H, Giaquinto S. Return to work following traumatic brain
399 injury: Trends and challenges. *Disabil. Rehabil.* [Internet]. 2007 [cited 2017 Dec
400 3];29:1387–95. Available from: <http://www.ncbi.nlm.nih.gov/pubmed/17729084>
- 401 2. Nayar M, Vanderstay R, Siegert RJ, Turner-Stokes L. The UK Functional Assessment
402 Measure (UK FIM+FAM): Psychometric Evaluation in Patients Undergoing Specialist
403 Rehabilitation following a Stroke from the National UK Clinical Dataset. *PLoS One*
404 [Internet]. 2016 [cited 2017 Dec 3];11:e0147288. Available from:
405 <http://www.ncbi.nlm.nih.gov/pubmed/26824696>
- 406 3. Turner-Stokes L. BSRM Standards for Rehabilitation Services Mapped on to the
407 National Service Framework for Long-Term Conditions BSRM Standards for
408 Rehabilitation Services, Mapped on to the National Service Framework for Long-Term
409 Conditions. 2009 [cited 2017 Dec 16]; Available from:
410 <https://www.bsrn.org.uk/downloads/standardsmapping-final.pdf>
- 411 4. Turner-Stokes L, Nyein K, Turner-Stokes T, Gatehouse C. The UK FIM+FAM:
412 development and evaluation. *Clin. Rehabil.* [Internet]. 1999 [cited 2017 Dec
413 16];13:277–87. Available from: <http://www.ncbi.nlm.nih.gov/pubmed/10460115>
- 414 5. British Society of Rehabilitation Medicine. BSRM Standards for Rehabilitation
415 Services Mapped on to the National Service Framework for Long-Term Conditions
416 BSRM Standards for Rehabilitation Services, Mapped on to the National Service
417 Framework for Long-Term Conditions [Internet]. London: 2009 [cited 2017 Dec 26].
418 Available from: <https://www.bsrn.org.uk/downloads/standardsmapping-final.pdf>
- 419 6. Law J, Fielding B, Jackson D, Turner-Stokes L. The UK FIM+FAM Extended

- 420 Activities of Daily Living module: evaluation of scoring accuracy and reliability.
421 Disabil. Rehabil. [Internet]. 2009 [cited 2017 Dec 16];31:825–30. Available from:
422 <http://www.ncbi.nlm.nih.gov/pubmed/19037776>
- 423 7. Turner-Stokes L, Siegert RJ. A comprehensive psychometric evaluation of the UK
424 FIM + FAM. Disabil. Rehabil. [Internet]. 2013 [cited 2017 Dec 16];35:1885–95.
425 Available from: <http://www.ncbi.nlm.nih.gov/pubmed/23384240>
- 426 8. Hawley CA, Taylor R, Hellowell DJ, Pentland B. Use of the functional assessment
427 measure (FIM+FAM) in head injury rehabilitation: a psychometric analysis. J. Neurol.
428 Neurosurg. Psychiatry [Internet]. 1999 [cited 2017 Oct 20];67:749–54. Available from:
429 <http://www.ncbi.nlm.nih.gov/pubmed/10567491>
- 430 9. UK FIM+FAM Users Group. The UK FIM+FAM (Functional Assessment Measure)
431 Version 2.2. [Internet]. Middlesex: 2010 [cited 2017 Oct 27]. Available from:
432 [https://www.kcl.ac.uk/nursing/departments/cicelysaunders/resources/FIMFAM-](https://www.kcl.ac.uk/nursing/departments/cicelysaunders/resources/FIMFAM-manual-v2.2-Sept-2012-print-double-sided.pdf)
433 [manual-v2.2-Sept-2012-print-double-sided.pdf](https://www.kcl.ac.uk/nursing/departments/cicelysaunders/resources/FIMFAM-manual-v2.2-Sept-2012-print-double-sided.pdf)
- 434 10. Hall KM, Bushnik T, Laskisic-Kazazic B, Wright J, Cantagallo A. Assessing traumatic
435 brain injury outcome measures for long-term follow-up of community-based
436 individuals. Arch. Phys. Med. Rehabil. [Internet]. 2001 [cited 2017 Oct 20];82:367–
437 74. Available from:
438 <http://www.sciencedirect.com/science/article/pii/S0003999301477380>
- 439 11. Nayar M, Vanderstay R, Siegert RJ, Turner-Stokes L, Burn J, Wood D. The UK
440 Functional Assessment Measure (UK FIM+FAM): Psychometric Evaluation in
441 Patients Undergoing Specialist Rehabilitation following a Stroke from the National
442 UK Clinical Dataset. PLoS One [Internet]. 2016 [cited 2017 Oct 18];11:e0147288.
443 Available from: <http://dx.plos.org/10.1371/journal.pone.0147288>
- 444 12. Reise SP, Moore TM, Haviland MG. Bifactor models and rotations: exploring the

- 445 extent to which multidimensional data yield univocal scale scores. *J. Pers. Assess.*
446 [Internet]. 2010 [cited 2017 Sep 30];92:544–59. Available from:
447 <http://www.ncbi.nlm.nih.gov/pubmed/20954056>
- 448 13. Ween JE, Alexander MP, D’Esposito M, Roberts M. Factors predictive of stroke
449 outcome in a rehabilitation setting. *Neurology* [Internet]. 1996 [cited 2017 Oct
450 20];47:388–92. Available from: <http://www.ncbi.nlm.nih.gov/pubmed/8757009>
- 451 14. Power T, Catroppa C, Coleman L, Ditchfield M, Anderson V. Do lesion site and
452 severity predict deficits in attentional control after preschool traumatic brain injury
453 (TBI)? *Brain Inj.* [Internet]. 2007 [cited 2017 Oct 20];21:279–92. Available from:
454 <http://www.tandfonline.com/doi/full/10.1080/02699050701253095>
- 455 15. NHS England. D02/S/a NHS standard contract for specialised rehabilitation for
456 patients with highly complex needs (all ages) [Internet]. 2013 [cited 2017 Dec 26].
457 Available from: [https://www.england.nhs.uk/wp-content/uploads/2014/04/d02-rehab-](https://www.england.nhs.uk/wp-content/uploads/2014/04/d02-rehab-pat-high-needs-0414.pdf)
458 [pat-high-needs-0414.pdf](https://www.england.nhs.uk/wp-content/uploads/2014/04/d02-rehab-pat-high-needs-0414.pdf)
- 459 16. Curran PJ, West SG, Finch JF. The robustness of test statistics to nonnormality and
460 specification error in confirmatory factor analysis. *Psychol. Methods* [Internet]. 1996
461 [cited 2016 May 2];1:16–29. Available from:
462 [http://apps.webofknowledge.com.ezproxy4.lib.le.ac.uk/full_record.do?product=UA&](http://apps.webofknowledge.com.ezproxy4.lib.le.ac.uk/full_record.do?product=UA&search_mode=GeneralSearch&qid=30&SID=S2WGEsVO85emPmGH2f7&page=1&doc=1)
463 [earch_mode=GeneralSearch&qid=30&SID=S2WGEsVO85emPmGH2f7&page=1&do](http://apps.webofknowledge.com.ezproxy4.lib.le.ac.uk/full_record.do?product=UA&search_mode=GeneralSearch&qid=30&SID=S2WGEsVO85emPmGH2f7&page=1&doc=1)
464 [c=1](http://apps.webofknowledge.com.ezproxy4.lib.le.ac.uk/full_record.do?product=UA&search_mode=GeneralSearch&qid=30&SID=S2WGEsVO85emPmGH2f7&page=1&doc=1)
- 465 17. George D, Mallery M. *SPSS for Windows Step by Step: A Simple Guide and*
466 *Reference, 17.0 update.* Boston, MA: Pearson; 2010.
- 467 18. West SG, Finch JF, Curran PJ. Structural equation models with non-normal variables:
468 Problems and remedies. In: H. HR, editor. *Structural equation modeling: Concepts,*
469 *issues, and applications.* Thousand Oaks, CA: Sage; 1995. p. 56–75.

- 470 19. Gorsuch RL, Hillsdale NJ. Factor Analysis. 2nd ed. New Jersey: Lawrence Erlbaum
471 Associates; 1983.
- 472 20. Zwick WR, Velicer WF. Comparison of five rules for determining the number of
473 components to retain. *Psychol. Bull.* 1986;99:432–42.
- 474 21. Fabrigar LR, Wegener DT, MacCallum RC, Strahan EJ. Evaluating the use of
475 exploratory factor analysis in psychological research. *Psychol. Methods* [Internet].
476 1999 [cited 2016 May 2];4:272–99. Available from:
477 http://apps.webofknowledge.com.ezproxy4.lib.le.ac.uk/full_record.do?product=UA&search_mode=GeneralSearch&qid=5&SID=S2WGEsVO85emPmGH2f7&page=1&doc=2
478
479 =2
- 480 22. Tabachnick BG, Fidell LS. Principal Components and Factor Analysis. In: Tabachnick
481 BG, Fidell LS, editors. *Using Multivariate Statistics*. Harlow: Pearson Education;
482 2014. p. 659–730.
- 483 23. Little TD, Lindenberger U, Nesselrode JR. On selecting indicators for multivariate
484 measurement and modeling with latent variables: When “good” indicators are bad and
485 “bad” indicators are good. *Psychol. Methods*. 1999;4:192–211.
- 486 24. Velicer WF, Fava JL. Affects of variable and subject sampling on factor pattern
487 recovery. *Psychol. Methods*. 1998;3:231–51.
- 488 25. Barrett P. Structural equation modelling: Adjudging model fit. *Pers. Individ. Dif.*
489 2007;42:815–24.
- 490 26. Hu L, Bentler PM. Cutoff criteria for fit indexes in covariance structure analysis:
491 Conventional criteria versus new alternatives. *Struct. Equ. Model. A Multidiscip. J.*
492 1999;6:1–55.
- 493 27. Kline RB. *Principles and practice of structural equation modeling*. 2nd ed. New York:
494 Guilford Press; 2005.

- 495 28. Bentler PM, Bonett DG. Significance tests and goodness-of-fit in the analysis of
496 covariance structures. *Psychol. Bull.* 1980;88:588–606.
- 497 29. Hu LT, Bentler PM. Cutoff criteria for fit indices in covariance structure analysis:
498 Conventional criteria versus new alternatives. *Struct. Equ. Model.* 1999;6 SRC-G:1–
499 55.
- 500 30. Cheung GW, Rensvold RB. Evaluating Goodness-of-Fit Indexes for Testing
501 Measurement Invariance. *Struct. Equ. Model. A Multidiscip. J.* 2002;9:233–55.
- 502 31. Jöreskog KG. How Large Can a Standardized Coefficient be? 1999 [cited 2017 Dec
503 16]; Available from:
504 [http://www.ssicentral.com/lisrel/techdocs/HowLargeCanaStandardizedCoefficientbe.p](http://www.ssicentral.com/lisrel/techdocs/HowLargeCanaStandardizedCoefficientbe.pdf)
505 [df](http://www.ssicentral.com/lisrel/techdocs/HowLargeCanaStandardizedCoefficientbe.pdf)
- 506 32. Cohen J. A power primer. *Psychol. Bull.* 1992;112:155–9.
- 507 33. Turner-Stokes L, Siegert RJ. A comprehensive psychometric evaluation of the UK
508 FIM + FAM. *Disabil. Rehabil.* 2013;35:1885–95.
- 509 34. Ponsford J, Draper K, Schibberger M. Functional outcome 10 years after traumatic
510 brain injury: Its relationship with demographic, injury severity, and cognitive and
511 emotional status. *J. Int. Neuropsychol. Soc.* [Internet]. 2008 [cited 2017 Dec
512 16];14:233–42. Available from: <http://www.ncbi.nlm.nih.gov/pubmed/18282321>
- 513 35. Gladman JRF, Harwood DMJ, Barer DH. Predicting the outcome of acute stroke:
514 prospective evaluation of five multivariate models and comparison with simple
515 methods. *J. Neurol. Neurosurgery, Psychiatry* [Internet]. 1992 [cited 2017 Dec
516 3];55:347–51. Available from: <http://jnnp.bmj.com/content/jnnp/55/5/347.full.pdf>
- 517 36. Tabachnick BG, Fidell LS, Height MA. *Using Multivariate* Needham Allyn &
518 Bacon. 2001;

- 519 37. Jongbloed, L. (1986). Prediction of function after stroke: A critical review. *Stroke*,
520 17(4), 765-776. doi: 10.1161/01.STR.17.4.765

521

522

523

524 Table 1

525 *Exploratory Factor Analysis (Maximum Likelihood Extraction with Promax Rotation) of the*526 *FIM+FAM items.**

	Focal-proximate		Diffuse-proximate	
	(n = 210)		(n = 208)	
	1	2	1	2
Eating	0.530	0.324	0.719	0.224
Swallowing	0.315	0.461	0.587	0.316
Grooming	0.611	0.349	0.676	0.313
Bathing	0.785	0.173	0.760	0.215
Dressing (upper)	0.758	0.191	0.859	0.115
Dressing (lower)	0.891	0.054	0.918	0.038
Toileting	0.883	0.032	0.955	0.008
Bladder (assist)	0.786	0.095	0.841	0.088
Bowel (assist)	0.802	0.116	0.884	0.080
Bed chair (transfer)	1.044	-0.121	1.065	-0.120
Toilet (transfer)	1.041	-0.129	1.055	-0.094
Tub/shower (transfer)	1.007	-0.174	1.016	-0.076
Car (transfer)	0.720	0.088	0.824	0.045
Locomotion (walking)	0.961	-0.088	0.908	0.013
Stairs	0.931	-0.148	0.946	-0.140
Community (mobility)	0.558	0.179	0.564	0.193
Comprehension	0.044	0.822	0.149	0.772
Expression	0.003	0.809	0.204	0.736
Reading	0.080	0.660	0.202	0.673

Writing	0.146	0.543	0.224	0.581
Speech intelligibility	0.091	0.575	0.303	0.555
Social interaction	-0.092	0.885	-0.035	0.904
Emotional status	-0.035	0.783	-0.066	0.862
Adjust to limits	-0.045	0.891	-0.117	0.992
Leisure activities	0.139	0.690	0.125	0.756
Problem solving	0.021	0.842	0.097	0.771
Memory	0.000	0.883	-0.096	0.961
Orientation	-0.155	0.976	-0.126	1.017
Concentration	-0.057	0.898	0.068	0.842
Safety	0.052	0.782	0.102	0.747

527 *Loadings that could be considered above 0.45 (“Fair”)³⁶ are bolded.

528

529 Table 2

530 *Confirmatory Factor Analysis Fit Statistics for the Different Models Proposed for FIM+FAM*

531 *Scale.*

	χ^2	df	$P = <$	CMIN	CFI	NNFI	RMSEA
	/DF						
Focal-proximate Sample ($n = 210$)							
2-factor	2544.29	404	.000	6.30	.764	.746	.159
3-factor (Nayar)	2090.74	374	.000	5.59	.803	.786	.148
4-factor (Turner-Stokes)	2356.47	400	.000	5.89	.784	.765	.153
Unidimensional	3809.62	405	.000	9.41	.624	.596	.201
2-factor (Bi)	2331.38	375	.000	6.22	.784	.749	.158
3-factor Bi (Nayar)	1198.53	345	.000	3.47	.902	.885	.109
4-factor Bi (Turner-Stokes)	1423.04	370	.000	3.84	.884	.863	.117
Diffuse-proximate Sample ($n = 207$)							
2-factor	3006.18	404	.000	7.44	.756	.737	.177
3-factor (Nayar)	2352.97	374	.000	6.29	.807	.790	.160
4-factor (Turner-Stokes)	2600.77	400	.000	6.50	.794	.776	.163
Unidimensional	4142.10	405	.000	10.23	.650	.624	.212
2-factor (Bi)	2690.70	375	.000	7.18	.783	.748	.173
3-factor Bi (Nayar)	1481.27	345	.000	4.29	.889	.870	.126
4-factor Bi (Turner-Stokes)	1713.52	370	.000	4.63	.874	.852	.133

532

533

534