

Empirical assessment of effect of publication bias on meta-analyses

A J Sutton, S J Duval, R L Tweedie, K R Abrams and D R Jones

BMJ 2000;320;1574-1577 doi:10.1136/bmj.320.7249.1574

Updated information and services can be found at: http://bmj.com/cgi/content/full/320/7249/1574

These include:

Data supplement	"Table and figures" http://bmj.com/cgi/content/full/320/7249/1574/DC1
References	This article cites 7 articles, 2 of which can be accessed free at: http://bmj.com/cgi/content/full/320/7249/1574#BIBL
	38 online articles that cite this article can be accessed at: http://bmj.com/cgi/content/full/320/7249/1574#otherarticles
Rapid responses	2 rapid responses have been posted to this article, which you can access for free at: http://bmj.com/cgi/content/full/320/7249/1574#responses
	You can respond to this article at: http://bmj.com/cgi/eletter-submit/320/7249/1574
Email alerting service	Receive free email alerts when new articles cite this article - sign up in the box at the top right corner of the article
Topic collections	Articles on similar topics can be found in the following collections Systematic reviews (incl meta-analyses): descriptions (132 articles) Other Journalology (371 articles)

Notes

To order reprints of this article go to: http://www.bmjjournals.com/cgi/reprintform Contributors: DAA and JME were instrumental in formulating the idea for the survey, and this was further developed in collaboration with IGD. All authors contributed to its design, the preparation of the questionnaire, and the identification of consultant psychiatrists in Scotland. SK and NMG carried out the pilot study, took responsibility for data entry and analysis, and advised on the interpretation of the data. DAA and JME were responsible for initial drafting of the paper, but all authors contributed to the final version. DAA and JME are the guarantors.

Funding: Grampian Healthcare NHS Trust. Competing interests: None declared.

- 1 Royal College of General Practitioners. Stress management in general practice. London: RCGP, 1993. (Occasional paper 61.)
- 2 Firth-Cozens J, Morrison LA. Sources of stress and ways of coping in junior house officers. *Stress Medicine* 1989;5:121-6.
- 3 Ramirez AJ, Graham J, Richards MA, Cull A, Gregory WM. Mental health of hospital consultants: the effects of stress and satisfaction at work. *Laneet* 1996;347:724-8.
- 4 Deary IJ, Agius RM, Sadler A. Personality and stress in consultant psychiatrists. *Int J Soc Psychiatry* 1996;42:112-23.
- 5 Chemtob CM, Hamada RS, Bauer G, Kinney B, Torigoe RY. Patients' suicide: frequency and impact on psychiatrists. Am J Psychiatry 1988; 145:224-8.
- 6 Litman RE. When patients commit suicide. Am J Psychiatry 1965;19: 570-84.
- 7 Menninger WW. Patient suicide and its impact on the psychotherapist. Bull Menninger Clin 1991;55:216-27.

- 8 Gralnick A. Suicide in the psychiatric hospital. *Child Psychiatry Hum Dev* 1993;24:3-12.
- 9 Kaye NS, Soreft SM. The psychiatrist's role, responses, and responsibilities when a patient commits suicide. Am J Psychiatry 1991;148:739-43.
- 10 O'Reilly RL, Truant GS, Donaldson L. Psychiatrists' experience of suicide in their patients. *Psychiatry J Univ Ottawa* 1990;15:173-6.
- 11 Little JD. Staff response to in-patient and out-patient suicide: what happened and what do we do? Aust NZJ Psychiatry 1992;26:162-7.
- 12 Goldstein LS, Buongiorno PA. Psychotherapists as suicide survivors. Am J Psychiatry 1984;38:392-8.
- 13 Bartels SJ. The aftermath of suicide on the psychiatric in-patient unit. Gen Hosp Psychiatry 1987;9:189-97.
- 14 Reder P, Duncan S. Reflections on child abuse inquiries. In: Peay J, ed. *Inquiries after homicide*. London: Duckworth, 1996.
- 15 Morris M. The aftermath of suicide. BrJ Nurs 1995;4:205-8.
- 16 Kendell RE, Pearce A. Consultant psychiatrists who retired prematurely in 1995-1996. Psychiatr Bull 1997;21:741-5.
- Eastman N. Inquiry into homicides by psychiatric patients: systematic audit should replace mandatory inquiries. *BMJ* 1996;313:1069-71.
 Schneidman ES. The management of the pre-suicidal, suicidal, and post-
- suited name and the second state of the pre-suited at suited at a postsuited patient. Ann Intern Med 1971;75:441-58.
- 19 Dewar I, Eagles JM, Klein S, Gray N, Alexander DA. Psychiatric trainees' experience of, and reactions to, patient suicide. *Psychiatr Bull* 1999;24:20-3.
- 20 Alexander DA. Human reactions to trauma. In: Greaves I, Porter KM, eds. Pre-hospital medicine. Principles and practice of immediate care. London: Arnold, 1999.

(Accepted 24 February 2000)

Empirical assessment of effect of publication bias on meta-analyses

A J Sutton, S J Duval, R L Tweedie, K R Abrams, D R Jones

Department of Epidemiology and Public Health, University of Leicester, Leicester LE1 6TP A J Sutton lecturer in medical statistics K R Abrams senior lecturer in medical statistics D R Jones professor in medical statistics

Division of Epidemiology, School of Public Health, University of Minnesota, Minneapolis, USA S J Duval assistant professor continued over

BMJ 2000;320:1574–7



Figures illustrating the method and funnel plots of all trials and a full table of results are available on the BMJ's website

Abstract

Objective To assess the effect of publication bias on the results and conclusions of systematic reviews and meta-analyses.

Design Analysis of published meta-analyses by trim and fill method.

Studies 48 reviews in Cochrane Database of Systematic Reviews that considered a binary endpoint and contained 10 or more individual studies. **Main outcome measures** Number of reviews with missing studies and effect on conclusions of meta-analyses.

Results The trim and fill fixed effects analysis method estimated that 26 (54%) of reviews had missing studies and in 10 the number missing was significant. The corresponding figures with a random effects model were 23 (48%) and eight. In four cases, statistical inferences regarding the effect of the intervention were changed after the overall estimate for publication bias was adjusted for.

Conclusions Publication or related biases were common within the sample of meta-analyses assessed. In most cases these biases did not affect the conclusions. Nevertheless, researchers should check routinely whether conclusions of systematic reviews are robust to possible non-random selection mechanisms.

Introduction

Selection bias is known to occur in meta-analyses because studies with results that are significant, interesting, from large well-funded studies, or of higher quality are more likely to be submitted, published, or published more rapidly than work without such characteristics.¹ A meta-analysis based on a literature search will thus include such studies differentially, and the resulting bias may invalidate the conclusions.

The best way to deal with these problems, which we shall collectively label "publication bias," is to avoid them. Recently, for example, a trial amnesty was announced that encouraged researchers to submit for publication reports of previously unpublished trials.² Additionally, steps are being taken to encourage the prospective registration of trials through trial registries.³ Although these steps may reduce the problem of publication bias in the future, it will remain a serious problem, and one that meta-analysts need to address for some time to come.

The simplest and most commonly used method to detect publication bias is an informal examination of a funnel plot.⁴ Formal tests for publication bias, such as those developed by Begg and Mazumdar⁵ and Egger et al,⁶ exist, but in practice few meta-analyses have assessed or adjusted for the presence of publication bias. A recent assessment of the quality of systematic reviews reported that only 6.5% and 3.2% of studies in high impact general and specialist journals respectively reported that a funnel plot had been examined.⁷ The uptake of any formal methods is lower still, although there are notable exceptions.⁸

The main aim of this paper is to assess what effect publication bias could have on the results and conclusions of meta-analyses of randomised controlled

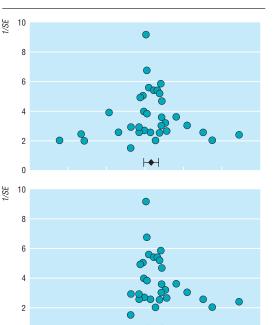


Fig 1 Typical funnel plot generated from 35 simulated studies (top) and same data with five missing studies showing a typical manifestation of publication bias (bottom)

0.5

1.5

Effect size

-0.5

trials in general. We applied the trim and fill method to a set of meta-analyses contained within the *Cochrane Database of Systematic Reviews*⁹ and estimated the numbers of missing trials and their effects on the inferences in these meta-analyses. This method both tests for the presence of publication bias and adjusts for it. It is simpler to implement than previously described methods,¹⁰ and simulation studies suggest that it may outperform more sophisticated methods in many situations.

Methods

0

-1.5

-1

A funnel plot is a plot of each trial's effect size against some measure of its size, such as the precision (used here), the overall sample size, or the standard error (fig 1, top).⁴ These plots are referred to as funnel plots because they should be shaped like a funnel if no publication bias is present. This shape is expected because trials of smaller size (which are more numerous) have increasingly large variation in the estimates of their effect size as random variation becomes increasingly influential. However, since smaller or non-significant studies are less likely to be published, trials in the bottom left hand corner (when a desirable outcome is being considered) of the plot are often omitted, creating a degree of asymmetry in the funnel (fig 1, bottom).

We used the "trim and fill" method to evaluate bias in funnel plots.¹⁰⁻¹² Firstly, the number of "asymmetric" trials on the right side of the funnel is estimated: these can broadly be thought of as trials which have no left side counterpart. These trials are then removed, or "trimmed," from the funnel, leaving a symmetric remainder from which the true centre of the funnel is estimated by standard meta-analysis procedures. The trimmed trials are then replaced and their missing counterparts imputed or "filled": these are mirror images of the trimmed trials with the mirror axis placed at the pooled estimate (see *BMJ*'s website for figure). This then allows an adjusted overall confidence interval to be calculated. A test for the presence of publication bias has also been derived from this method, based on the estimated number of missing trials.¹¹

Selection and analysis of studies

We examined all reviews contained in the Cochrane Database of Systematic Reviews (1998, issue 3).9 Reviews including 10 or more trials reporting a binary outcome measure were included in the assessment. At most, one meta-analysis from each review was included, and when more than one met the inclusion criteria, the one containing the most studies was selected. If two or more contained the same number of studies, the one listed first in the review was chosen. The (log) odds ratio measure was used for analysis, and where data were sparse, a continuity correction of 0.5 was used.¹³ For consistency, funnel plots were reflected about zero for meta-analyses in which the reported outcome measure was "undesirable" so that the left hand side of the funnel plot was scrutinised for publication bias in every case.

For each meta-analysis, we trimmed the trials considered to be symmetrically unmatched (assumed to be the *k* right-most trials in each funnel). The number *k* was estimated by an iterative procedure (details given elsewhere¹⁰⁻¹²) using the estimator denoted L_0 .^{11 12} We used fixed and random effects models to estimate the overall effect in order to assess the impact of choice of model on publication bias. The estimated effect of the missing trials provides an indication of whether the imputed missing studies affect the overall result of the meta-analysis.

Results

Included studies

At the time of the investigation, the *Cochrane Library* contained 397 reviews that included a meta-analysis. Of these, 49 included 10 or more trials with at least one dichotomous outcome. However, one of these did not assess a comparative effect and was excluded, leaving 48 meta-analyses for assessment. The number of trials included in each dataset ranged from 10 to 47 (median 13), with only four analyses including more than 20 trials.

Random effects meta-analyses of these datasets, ignoring possible publication bias, resulted in 28 estimates that reached significance at the 5% level, with the remaining 20 being inconclusive. The corresponding figures were 30 and 18 with a fixed effects model.

The funnel plots for all 48 meta-analyses are available on the *BMJ*'s website. Each trial's log odds ratio is plotted against the reciprocal of its standard error. The distribution of individual trial effect sizes is often highly irregular. Only a few of the plots closely conform to the classic funnel shape.

Estimated numbers of missing studies

Details of the estimates of the number of trials missed because of publication bias for each review are Department of Biostatistics, School of Public Health, University of Minnesota R L Tweedie *head*

Correspondence to: A J Sutton ajs22@le.ac.uk

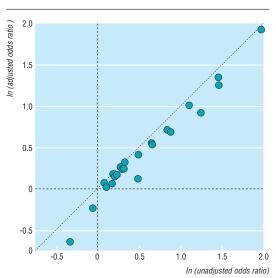


Fig 2 Scatter plot of unadjusted versus adjusted odds ratios after filling with trim and fill method under random effects model. Dotted line is line of equality, and vertical distance away from this line indicates magnitude by which original pooled estimates have been reduced by adjusting for publication bias

available on the *BMJ*'s website. In all, 23 meta-analyses were estimated to have some degree of publication bias $(L_0 > 0)$ with the random effects model; this increased to 26 with the fixed effects model. The number of missing trials was significant if $L_0 > 3$ for the range of trials included in the meta-analyses evaluated here.¹² Eight meta-analyses reached this critical level under the random effects model. These estimates suggest that about half of meta-analyses may be subject to some level of publication bias and about a fifth have a strong indication of missing trials.

Changes in significance and magnitude of the overall pooled estimates

Imputing missing trials changed the estimates of the overall effect for all meta-analyses in which one or more trials was estimated as missing. Use of the random effects model indicated that in four reviews this would lead to significant changes in the conclusions. Three meta-analyses that were considered significant at the 5% level in the original analysis became non-significant (studies 5, 10, 13), and one that was considered non-significant became significantly negative (study 30). These results seem plausible in the light of the trial distribution in figure 2, and thus we deduce that around 5-10% of meta-analyses may be interpreted incorrectly because of publication bias.

Figure 2 shows the effect of adjustment by the trim and fill method on the odds ratios for the 23 meta-analyses estimated to have one or more missing trials. Although only modest changes in outcome are observed in most instances, in six studies with a positive effect (studies 1, 3, 5, 10, 24, and 47) the reduction is greater than 30%. In two cases (studies 30 and 48) with a negative effect, adding missing studies had an even stronger effect (see table on *BMJ*'s website for details).

Discussion

We found a clear indication of publication bias within this sample of studies from the Cochrane Database of Systematic Reviews. About half had some indication of publication bias, with a fifth having a strong indication. The potential effect of such bias varies from dataset to dataset. In some instances (for example, study 3) the number of studies estimated to be missing was relatively large but the adjusted estimate of effect size was only slightly lower than the original estimate. This can largely be explained by the presence of one or more studies that are much larger than those filled. These large studies have sufficient influence over the overall pooled estimate that asymmetry in the lower part of the funnel makes little impression on the overall pooled result. In other cases (for example, study 5), the estimated number of missing studies is smaller but the review's conclusions are altered.

The only other assessment of the impact of funnel plot asymmetry on a general collection of metaanalyses which is known to us is that of Egger et al,⁶ who applied their test to an earlier edition of the *Cochrane Database of Systematic Reviews*. Their inclusion criteria were less stringent than ours, including meta-analyses of categorical outcomes with five or more trials (compared with our minimum requirement of 10). They found significant indication of bias in five meta-analyses out of 38 examined (P < 0.1). Our decision to include only meta-analyses with a minimum of 10 trials was largely arbitrary; however, five studies are usually too few to detect an asymmetric funnel.

Use of bias assessment

Although the development of methods for assessing publication bias has a reasonably long history, these methods are rarely used in practice.⁷ We searched the 48 original review reports from which the 48 meta-analyses were taken for references to publication bias and descriptions of any steps taken to deal with it. Thirty (63%) made no reference to publication bias. Five reviews mentioned examining a funnel plot, and three used the test of Egger et al. Since many of these reviews were written before that test was published, it has made a relatively high impact among reviewers subsequently. Generally, it was the more recently conducted reviews that had considered or tested for funnel plot asymmetry.

One possible reason for the lack of uptake of methods to deal with publication bias is that previous approaches have involved modelling methods that are difficult to implement and require lengthy calculations. The trim and fill method is both theoretically simple and easy and quick to implement. It may therefore be appropriate for routine use in evaluating metaanalyses. Our results seem plausible, with the four studies whose conclusions changed on adjustment all having visually skewed funnel plots.

Validity of results

Some cautionary remarks are needed in assessing our results. Since the method is based on the lack of symmetry in the funnel plot, and asymmetry might be due to factors other than publication bias,¹⁴ the results produced by trim and fill may not always reflect correc-

tion for publication bias. Moreover, the odds ratio outcome was used exclusively in this investigation. The appearance of a funnel plot can depend on the outcome measure used, and different results might be obtained in some instances if the risk difference or relative risk scale is used. The sensitivity of assessments of publication bias to the outcome measure used requires further investigation.

The idea of adjusting the results of meta-analyses for publication bias and imputing "fictional" studies into a meta-analysis is controversial.¹⁵ We certainly would not rely on results of imputed studies in forming a final conclusion, partly because asymmetry in a funnel plot may be due to factors other than publication bias. Any adjustment method should be used primarily as a form of sensitivity analysis, to assess the potential effect of missing studies on the meta-analysis, rather than as a means of adjusting results themselves.

If, as our study indicates, missing studies change the conclusions in less than 10% of meta-analyses, publication bias, although widespread, may not be a major practical problem. On the other hand, the fact that almost half the funnel plots examined seemed to exhibit some asymmetry leads us to conclude that routine evaluation for this bias should be an important step in any systematic review.

Contributors: AJS had the original idea for the study, extracted the relevant data from the Cochrane Database of Systematic Reviews, and assisted with the interpretation of the results. SJD assisted with the study design, carried out the trim and fill analyses, and assisted with the interpretation of the results. RLT assisted with the study design and the interpretation of the results. KRA and DRJ assisted with the interpretation of the results. The paper was written and revised jointly by all authors. AJS is the guarantor for this paper.

Funding: None. Competing interests: None declared.

- 1 Song F, Easterwood A, Gilbody S, Duley L, Sutton, AJ. Publication bias. In: Stevens A, Abrams K, Brazier J, Fitzpatrick R, Lilford R, eds. Handbook of research methods for evidence-based health care-insights from the NHS HTA programme. London: Sage Publications (in press).
- Horton R. Medical editors trial amnesty. Lancet 1997;350:756
- Easterbrook PJ. Directory of registries of clinical trials. Stat Med 1992;11:345-423.

What is already known on this topic

Meta-analyses are subject to bias because smaller or non-significant studies are less likely to be published

Most meta-analyses do not consider the effect of publication bias on their results

What this study adds

A simple trim and fill adjustment method on studies in the Cochrane database suggests that publication bias may be present to some degree in about 50% of meta-analyses and strongly indicated in about 20%

Publication bias affected the results in less than 10% of meta-analyses

Researchers should always check for the presence of publication bias and perform a sensitivity analysis to assess the potential impact of missing studies

- Begg CB. Publication bias. In: Cooper H, Hedges LV, eds. The handbook of research synthesis. New York: Russell Sage Foundation, 1994:399-409. 4
- Begg CB, Mazumdar M. Operating characteristics of a rank correlation test for publication bias. *Biometrics* 1994;50:1088-101. Egger M, Smith GD, Schneider M, Minder C. Bias in meta-analysis
- $\mathbf{6}$ detected by a simple, graphical test. BMJ 1997;315:629-34
- 7 Tallon D, Schneider M, Egger M. Quality of systematic reviews published in high impact general and specialist journals. Proceedings of 2nd symposium on systematic reviews: beyond the basics, Oxford, 1999. (www.ihs.ox.ac.uk/csm/talks.html#p8; accessed 1 June 2000.) Linde K, Clausius N, Ramirez G, Melchart D, Eitel F, Hedges LV, et al. Are
- 8 the clinical effects of homoeopathy placebo effects? A meta-analysis of placebo-controlled trials. *Lancet* 1997;350:834-43. Cochrane database of systematic reviews. In: Cochrane Collaboration.
- Cochrane Library. Issue 3. Oxford: Update Software, 1998 10 Duval S, Tweedie R. Practical estimates of the effect of publication bias in
- meta-analysis. *Australasian Epidemiologist* 1998;5:14-7. 11 Duval S, Tweedie R. A non-parametric "trim and fill" method of assessing
- publication bias in meta-analysis. *J Am Stat Ass* (in press). 12 Duval S, Tweedie R. Trim and fill: a simple funnel plot based method of
- testing and adjusting for publication bias in meta-analysis. *Biometrics* (in ress)
- 13 Sankey SS, Weissfeld LA, Fine MJ, Kapoor W. An assessment of the use of the continuity correction for sparse data in meta-analysis. Con In Statistics–Simulation and Computation 1996;25:1031-56. 14 Petticrew M, Gilbody S, Sheldon TA. Relation between hostility and cor-
- onary heart disease: evidence does not support link. BMJ 1999;319:917.
- 15 Begg CB. Comment on Givens GH, Smith DD, Tweedie RL. Publication bias in meta-analysis: a Bayesian data-augmentation approach to account for issues exemplified in the passive smoking debate. Stat Sci 1997;12:221-50.

(Accepted 2 March 2000)

Back pain in Britain: comparison of two prevalence surveys at an interval of 10 years

Keith T Palmer, Kevin Walsh, Holly Bendall, Cyrus Cooper, David Coggon

In Britain, as in many other countries, back pain is a major cause of disability, especially in adults of working age. During the decade to 1993, outpatient attendances for back pain rose fivefold, and the number of days of incapacity from back disorders for which social security benefits were paid more than doubled.¹ It is unclear whether this represents an increase in the occurrence of diseases affecting the back or a change in people's behaviour when they have symptoms. To address this question we compared the prevalence of low back pain and associated disability in two postal surveys 10 years apart.

Subjects, methods, and results

Both surveys were approved by the relevant local ethics committees. The first was conducted during 1987-8 and obtained information from 2667 men and women randomly selected from the lists of 136 general practitioners in eight geographically dispersed locations in Britain (59% response rate).² Of these, 2596 were aged 20-59 years at the time of completing the questionnaire. The investigation focused on occupational and other risk factors for back symptoms and included a question about the occurrence of back pain that had

Editorial by Croft Correspondence to: D Coggon dnc@mrc.soton.ac.uk continued over

BMI 2000:320:1577-8