

Title Page

Title: The backwards comparability of wrist worn GENEActiv and waist worn ActiGraph accelerometer estimates of sedentary time in children

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Abstract

Objectives: To examine the backward comparability of a range of wrist-worn accelerometer estimates of sedentary time (ST) with ActiGraph 100 count·min⁻¹ waist ST estimates.

Design: Cross-sectional, secondary data analysis

Method: One hundred and eight 10-11-year-old children (65 girls) wore an ActiGraph GT3X+ accelerometer (AG) on their waist and a GENEActiv accelerometer (GA) on their non-dominant wrist for seven days. GA ST data were classified using a range of thresholds from 23-56 mg. ST estimates were compared to AG ST 100 count·min⁻¹ data. Agreement between the AG and GA thresholds was examined using Cronbach's alpha, intraclass correlation coefficients (ICC), limits of agreement (LOA), Kappa values, percent agreement, mean absolute percent error (MAPE) and equivalency analysis.

Results: Mean AG total ST was 492.4 minutes over the measurement period. Kappa values ranged from 0.31-0.39. Percent agreement ranged from 68-69.9%. Cronbach's alpha values ranged from 0.88-0.93. ICCs ranged from 0.59-0.86. LOA were wide for all comparisons. Only the 34 mg threshold produced estimates that were equivalent at the group level to the AG ST 100 count·min⁻¹ data though sensitivity and specificity values of ~64% and ~74% respectively were observed.

Conclusions: Wrist-based estimates of ST generated using the 34 mg threshold are comparable with those derived from the AG waist mounted 100 count·min⁻¹ threshold at the group level. The 34 mg threshold could be applied to allow group-level comparisons of ST with evidence generated using the ActiGraph 100 count·min⁻¹ method though it is important to consider the observed sensitivity and specificity results when interpreting findings.

Keywords: accelerometry, physical activity, sedentary behaviour, children, raw acceleration signals, measurement

Introduction

Sedentary behaviour (SB) has received increased attention across recent years as a behaviour that may detrimentally affect children's health. Whether SB influences health independent of physical activity (PA) is deemed to be a controversial topic, with some studies demonstrating the negative effects of reallocating moderate-to vigorous PA (MVPA) to SB^{1,2}, and others reporting limited evidence that SB is associated with health independent of MVPA³. Nonetheless, researchers are interested in measuring youth movement behaviours including SB to explore health associations, investigate secular trends, and establish intervention effects.

Sedentary behaviour is defined as any waking behaviour characterised by an energy expenditure of ≤ 1.5 metabolic equivalents (METs) while in a sitting, reclining or lying posture⁴. Despite the SB definition referring to posture, many researchers use accelerometers to classify SB as an absence of, or little, registered dynamic acceleration⁵. While widely used, it should be noted that this approach does not consider posture.

Historically children's SB or sedentary time (ST: the time spent for any duration or in any context in sedentary behaviours⁴) was assessed using waist-worn ActiGraph accelerometers with a threshold of ≤ 100 vertical axis count \cdot min⁻¹ used as the upper boundary for ST. This approach has demonstrated acceptable agreement with measures that classify posture such as the activPAL⁶ and those that provide an estimate of energy expenditure, for example indirect calorimetry⁷. Recently the field has moved towards that of wrist accelerometry due to superior wear compliance⁸. Despite observing better compliance, and moderate-to-strong correlations between acceleration data collected using wrist GENEActiv and waist ActiGraph accelerometer placements⁹, wrist placements generally result in higher estimates of physical activity, therefore wrist and hip data are not directly comparable without correcting for these differences⁸. Researchers have also begun to make use of raw acceleration signal analysis to remove the proprietary nature of counts-based data and improve comparability between different devices¹⁰. Although this is advantageous for studies moving forward, a wealth of SB data exists using the ≤ 100 count \cdot min⁻¹ threshold applied to data from hip-worn ActiGraphs (for example: ¹¹,

¹). Therefore, the ability to compare new raw acceleration derived estimates with those generated using the ≤ 100 count \cdot min⁻¹ method would be useful for researchers in the field.

In a previous study authors generated and cross-validated a GENEActiv (GA) wrist threshold of 51 mg with the intention of providing comparable estimates of ST to those generated using the traditional ActiGraph waist-worn ≤ 100 count \cdot min⁻¹ method. The open source R package GGIR was used to calculate average magnitude of dynamic acceleration, known as the Euclidean Norm Minus One (ENMO), from raw acceleration data¹², applying the newly generated (51 mg) and published thresholds for SB^{13, 14}. The comparability of ST estimates between the newly generated 51 mg threshold, the other empirical GA raw acceleration thresholds, and ActiGraph waist-worn ≤ 100 count \cdot min⁻¹ data was examined. Results demonstrated a lack of equivalence for the 51 mg threshold and existing GENEActiv empirical wrist ST thresholds^{15, 16}. The study provided some preliminary evidence of group-level agreement for the 36 mg empirical threshold¹³ which was originally intended to classify wrist-worn ActiGraph data. A study limitation, however, was that the individual level agreement was undetermined. At the study conclusion, the authors called for the backwards compatibility of ST estimates to be examined further, both at the individual *and* group levels covering a broad range of ENMO thresholds between the lowest (23 mg) and highest (56 mg) thresholds used to date by researchers in the field. This would enable researchers to establish the most comparable threshold to use when comparing to earlier estimates of ST from ActiGraph data. The lack of evidence related to the comparability of ST estimates currently presents a challenge for researchers when attempting to compare data to those previously recorded using hip and count methods. More investigation is required, therefore, to confirm whether the 36 mg proposed in the previous study represents the optimal threshold to use for this purpose.

The backwards comparability of wrist generated ENMO assessed MVPA with traditional accelerometer counts-based data using a range of waist-worn ActiGraph MVPA thresholds has been recently demonstrated. The study proposed wrist ENMO thresholds that gave estimates of MVPA that are comparable with waist measured counts-based data classified using empirical ActiGraph thresholds¹⁷.

To date, the backwards comparability for the range of ENMO-derived sedentary behaviour/time estimates has not been comprehensively examined. The aim of this secondary data analysis was therefore to extend previous work by using a wide range of both empirical and arbitrary ST thresholds to examine the backwards comparability of wrist-worn accelerometer estimates of ST with waist-worn ActiGraph 100 count·min⁻¹ ST. The study also develops previous work by investigating the extent of backwards compatibility at both the individual and group levels.

Methods

This is a secondary data analysis, and the methods for the study have been previously published elsewhere ^{8, 12, 15}. Briefly, after gaining institutional ethical approval, parental/carers consent and child assent, 10-11-year-old children were involved in this study (65 girls). Body mass was assessed to the nearest 0.1 kg (Seca, Birmingham, UK) and stature was assessed to the nearest 0.1 cm using a portable height meter (Leicester Height Measure, Seca, Birmingham, UK) during school-based data collection sessions conducted between January - May 2014.

Two tri-axial accelerometers (GENEActiv; Activinsights, Cambs, UK and ActiGraph GT3X+; ActiGraph, Pensacola, FL) were used to assess sedentary time. The GENEActiv (GA) was worn on the non-dominant wrist, and the ActiGraph (AG) worn on the right hip for waking hours over seven consecutive days. Participants were instructed to remove the monitors when engaging in water-based activities and also when sleeping. Both monitors were initialised to record at a sampling frequency of 100 Hz using the same computer.

Data generated by the AG devices were processed using ActiLife v 6.11.4 software (ActiGraph, Pensacola, FL). Consistent with previous research ^{15, 18}, for the AG devices, non-wear was defined as 20 minutes of consecutive zero counts (1-minute spike tolerance) and was subtracted from daily wear time. A valid day was defined as ≥ 540 min for weekdays ¹⁹ and ≥ 480 min for weekend days ²⁰. Consistent with a previous study ¹⁵, the valid weekend and weekdays with the longest wear for each participant were retained for analysis. Where participants did not have a valid weekend day, only their

longest valid weekday was retained for analysis. Data for the included days were converted into 1 second csv files, with non-wear times manually removed at a later step in data reduction described later. Sedentary time was defined as $\leq 100 \text{ count}\cdot\text{min}^{-1}$ ²¹ and coded accordingly. Data generated by the GA monitors were saved as binary files after being downloaded using GENEActiv v 2.2 software (Activinsights, Cambs, UK). GA data were processed in R using the GGIR package version 1.1-4 to calculate the ENMO-derived average magnitude of dynamic acceleration. ENMO is vulnerable to calibration errors, therefore to correct for sensor calibration errors, autocalibration was completed²². Non-wear for the GA data was scored using 60 minutes moving windows with 15 minutes increments and imputation was completed²³. ENMO values were expressed in average mg per 1 second epoch²³, and GA data for the corresponding AG week and/or weekend days were retained for further analysis.

Time stamps for the GA and AG were synchronised, and data were merged resulting in one csv file for each participant. Periods of non-wear were manually removed from both the AG and GA data according to the wear details generated by the ActiLife AG analysis. Therefore all epochs remaining in the dataset contained 'wear' data for both devices. After non-wear periods were removed, data were then reduced to 1-minute epochs and AG data were scored as sedentary or active using vertical axis $100 \text{ count}\cdot\text{min}^{-1}$ as the reference value for ST²¹. In the previous study ST thresholds of 23 mg (obtained by solving the Hildebrand et al., (2014) regression equation), 36 mg¹³, 51 mg (newly generated and cross-validated threshold), and 56 mg¹³ were used. In Step 1 of the analysis for the current study we extend these results to examine comparability of GA ST data classified using a wider range of thresholds. This included a recently published ST threshold of 52 mg that was generated by a child-specific calibration circuit²⁴ and arbitrary thresholds of 30 mg, 40 mg and 45 mg which were chosen to cover the range of thresholds. The final thresholds included in Step 1 were therefore: 23 mg, 30 mg, 36 mg, 40 mg, 45 mg, 52 mg and 56 mg. This approach resulted in a range of ST thresholds with which to compare to the AG vertical axis $100 \text{ count}\cdot\text{min}^{-1}$ reference.

Following calculation of descriptive statistics to describe the participant group, the GA ST estimates were compared to the AG $100 \text{ count}\cdot\text{min}^{-1}$ estimates at the group level by calculating Cronbach's alpha,

intraclass correlation coefficients (ICC), Kappa values and percent agreement. Individual level estimates were compared by calculating limits of agreement (LOA), correlations between bias and mean sedentary time (AG and GA) and mean absolute percent error (MAPE, %). Null hypothesis testing is not appropriate when considering the comparability between estimates ²⁵, therefore equivalency analysis was also performed to establish the equivalence of group level estimates of ST on average. A 95%equivalence test was completed to establish whether the 90% confidence intervals for the range of GA ST thresholds completely fell within the zone of equivalence, defined as $\pm 10\%$ of the mean AG 100 count \cdot min⁻¹ classified ST.

In Step 2 of the analysis, results from step 1 were used to identify the likely range of most comparable thresholds, and further thresholds within this range were then added to the analysis to attempt to find the optimal threshold. Analyses were also completed separately by sex to further examine the comparability of estimates. To provide a more stringent comparison, for the second step in analysis, the zone of equivalence for the group-level equivalence test was reduced to $\pm 5\%$ of the mean AG 100 count \cdot min⁻¹ classified ST. Analysis was completed using IBM SPSS Statistics v.24 (IBM, Armonk, NY) and Microsoft Excel 2016 (Microsoft, Redmond, WA).

Results

Participant characteristics, the number of days included in analyses, weekday and weekend day wear times for boys and girls have been published previously for this population ¹⁵, and are displayed in Table 1.

[TABLE 1 ABOUT HERE]

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176 Analysis Step 1. Table 2 summarises the comparisons between the $AG \leq 100 \text{ count} \cdot \text{min}^{-1}$ and various
177 GA ST estimates. Kappa values ranged from 0.31-0.39, representing ‘fair’ agreement ²⁶. Percent
178 agreement ranged from 68-69.9%. Cronbach’s alpha values were 0.88 for the 23 mg threshold,
179 suggesting a good level of consistency, where all other Cronbach’s alpha values were >0.9 , suggesting
180 excellent levels of consistency. ICCs ranged from 0.59 for the 23 mg threshold (moderate reliability) to
181 0.86 for the 36 mg threshold (good reliability). Supplementary content A displays the Bland-Altman
182 plots for the comparisons. LOA were wide for AG - GA comparisons, with the narrowest limits
183 observed for the AG v 36 mg comparison (lower LOA = -230.47 upper LOA = 194.81 minutes), with
184 systematic bias apparent. All thresholds from 36 mg and above showed negative bias illustrated by
185 mean bias and the correlation results (i.e. higher GA ST estimates than AG). The highest negative bias
186 observed for the 56 mg threshold. MAPE (%) ranged from 15.8% for the 36 mg threshold to 40.7% for
187 the 56 mg. The results of the equivalency analysis found that only the ST estimates generated by the
188 36 mg threshold could be considered statistically equivalent to the $AG \leq 100 \text{ count} \cdot \text{min}^{-1}$ on average at
189 the group level with 90% CI’s falling completely within the $\pm 10\%$ zone of equivalence. Thresholds ≤ 30
190 mg appeared to underestimate and ≥ 40 mg appeared to overestimate ST in comparison to the ST
191 reported using $AG \leq 100 \text{ count} \cdot \text{min}^{-1}$. Therefore, for analysis Step 2 thresholds of 34 mg and 35 mg
192 were included to examine the optimum threshold and analyses were repeated for the whole cohort and
193 separately by sex. The zone of equivalence was reduced to $\pm 5\%$ for the group-level equivalency
194 analysis. Table 3 summarises the comparisons between the $AG \leq 100 \text{ count} \cdot \text{min}^{-1}$ and 34 mg and 35 mg
195 GA ST estimates and includes sensitivity and specificity information. Mean bias was low for the 34 mg
196 threshold though wide limits of agreement were observed for both thresholds and MAPE% was similar
197 to that observed for the 36 mg threshold at 16.2% (34 mg) and 15.8% (35 mg). Sensitivity values (true
198 positive) were similar between the thresholds, at 63.6% and 64.8% for 34 mg and 35 mg respectively.
199 Specificity values (true negative) were also similar, at 74.2% and 73.4% for the 34 mg and 35 mg
200 thresholds respectively. Boys’ data displayed wider limits of agreement, higher MAPE% and slightly

higher sensitivity values for both thresholds in comparison to girls, though % agreement, Cronbach's alpha, ICC, Kappa and specificity values were similar. The results of the equivalence analysis for all threshold comparisons are displayed in Figure 1. Only the ST estimates generated by the 34 mg threshold were statistically equivalent to the AG ≤ 100 count·min⁻¹ on average at the group level with 90% CI's falling completely within the $\pm 5\%$ zone of equivalence.

[TABLE 2 ABOUT HERE]

[TABLE 3 ABOUT HERE]

[FIGURE 1 ABOUT HERE]

Discussion

A wealth of existing accelerometer data has used the threshold of ≤ 100 count·min⁻¹ applied to waist-worn ActiGraphs to determine time spent sedentary (for example large studies using the International Children's Accelerometry Database ^{11, 1}). As the discipline moves increasingly towards raw acceleration data processing and wrist-worn monitors, the ability for researchers to compare data between studies that have used counts-based processing methods and waist-worn monitors is important. The aim of this study was therefore to examine the backwards comparability of wrist-worn accelerometer estimates of sedentary time (ST) with ActiGraph 100 count·min⁻¹ waist ST estimates using a range of empirically determined and arbitrary raw acceleration thresholds for wrist-worn monitors.

This study has demonstrated moderate to excellent ICC values, and moderate to good Cronbach's alpha values at the group level for all the GA thresholds. Despite these results, Kappa values were 'fair' and large MAPE values (individual level) were observed. In addition, wide limits of agreement (individual level) were observed between all GA thresholds and the AG standard. Systematic bias was evident, indicating that as estimates of ST increased so did the bias. Equivalency analysis found no thresholds produced estimates of ST that could be considered statistically equivalent on average at the group level in comparison to the AG standard with the exception of the 34 mg threshold. The wide limits of agreement, MAPE and bias results, in the presence of high consistency, as evidence by high ICC and Cronbach's alpha, highlights the importance of considering a range of analyses at the individual and group levels when examining the comparability of ST estimates.

Our previous work called for studies to investigate the backwards compatibility of ST estimates ¹⁵. In the present study this was addressed by examining a broad range of thresholds, both empirically determined and arbitrary, and out of the selected thresholds the 34 mg threshold provided the ST estimates most comparable to the ActiGraph 100 count·min⁻¹ waist ST at the group level. Despite this, the limits of agreement showed that ST estimates generate using the 34 mg threshold ranged from $\pm \sim 3$ hrs in comparison to the AG estimates, therefore suggesting the 34 mg should not be used for individual level comparison. In our previous study we established that the 36 mg threshold, provided

equivalent estimates of wrist ST for the GENEActiv as the ≤ 100 count \cdot min⁻¹ standard at the group level when using a $\pm 10\%$ zone of equivalence. The present study that used a more stringent $\pm 5\%$ zone of equivalence suggests that 34 mg may provide a more accurate comparison at the group level, and furthermore suggests that lower and higher thresholds across the range are not appropriate for this purpose. ActiGraph accelerometers are known to produce lower ENMO values than GA devices ⁸, though recent evidence suggests the GA and AG devices provide equivalent estimates between the 30-50 mg range ¹⁶. Irrespective of potential differences between devices, at the group level the 34 mg threshold provided the most comparable estimates of ST to the AG hip ≤ 100 count \cdot min⁻¹ standard, so could be used for comparative purposes across studies moving forward.

Despite exhibiting group level equivalency, the sensitivity and specificity values suggest that for every 100 minutes of ST classified by the ActiGraph, the GA 34 mg threshold would classify ~64 minutes of ST. Therefore any comparisons between studies using the wrist worn 34 mg threshold and studies using the waist worn AG ≤ 100 count \cdot min⁻¹ method should bear the sensitivity and specificity results in mind when interpreting findings. Furthermore, the 5% zone of equivalence provides a range of ~50 minutes of sedentary time which the 34 mg estimates fell within. Whether a potential difference of $\pm \sim 50$ minutes is clinically meaningful or whether that would provide estimates that are sensitive to change is open to debate. Recent evidence suggests that the reallocation of 15 minutes of sedentary time to moderate to vigorous physical activity (MVPA) predicted changes in obesity and fitness outcomes in children². However, such evidence relies on sedentary time to be reallocated to MVPA, and the impact of reallocation of time to light intensity physical activity or stationary behaviours independent of MVPA remains unclear. Where group-level comparisons with data collected using the AG hip ≤ 100 count \cdot min⁻¹ standard are useful the 34 mg threshold can be applied, though where precise estimates of sedentary time or behaviour are required to demonstrate intervention effectiveness or individual level changes alternative methods may be required. Whether the estimates of ST from the GA and AG reflect actual ST remains open to debate. Indeed it is questionable whether the absence or low levels of acceleration should be used in isolation to classify ST, especially considering the postural component that is integral to the definition of sedentary behaviour. Examining the accuracy of measuring ST was not the aim of

the present study *per se*, and as such represents a different research question to be addressed in the future. There are, however, ways of processing accelerometer data to classify posture that do not require the use of additional devices or monitoring periods. One example is the sedentary sphere ²⁷, which classifies assumed postural changes based on acceleration signals, arm orientation and wrist orientation. Although this approach has shown promise in adult populations, it has not yet been validated in children and so its utility in this population has not been established. Nonetheless, estimates based on new approaches, irrespective of the method, still raises questions regarding the comparability with the large volume of existing literature therefore a pragmatic solution is warranted.

There are some limitations to the present study. We used a 1-minute epoch to determine time spent sedentary to allow a comparison to the AG hip ≤ 100 count·min⁻¹ standard. The majority of children's ST data using AG hip ≤ 100 count·min⁻¹ utilises 1-minute epochs, therefore this approach was necessary to address the study aims. It is well established that children's physical activity behaviours are sporadic in nature ^{28, 29} and though high frequency monitoring is required to detect movement at higher intensities, the 1 minute epoch is unlikely to influence recorded ST which is generally accrued in bouts lasting >2 minutes ³⁰. In addition, the group of participants involved in this study were all from the same geographical location in North-West England and a narrow age-range, therefore their ST behaviours may not be representative of different populations and groups. We included a maximum of 2 days of data (one weekend and one weekday) for each participant, therefore the sedentary levels of participants are not reflective of their habitual patterns. However, the volume of data included allows for comparison between devices and signal classification, where 7 day's data would have been prohibitive in terms of file size. Furthermore, a waking hours accelerometer protocol was used. Therefore recent studies using 24-hour protocols may require further investigation to examine the backward compatibility of data, including sleep classification in addition to ST estimates.

To the best of our knowledge, this is the first study to examine the backward comparability of wrist assessed sedentary time with ActiGraph 100 count·min⁻¹ waist ST estimates. The results of the study suggest that the 34 mg threshold produced the most comparable estimates of ST and could be used to

classify data for group-level comparison with previously published studies that used the 100 count·min⁻¹ threshold.

Conclusions

Despite observing high ICC and Cronbach's alpha values, the results suggest that the all but one of the wrist mounted, raw acceleration derived ST estimates should not be directly compared with those derived from the 100 count·min⁻¹ waist mounted AG threshold. The 34 mg threshold may provide comparable ST estimates at the group level, and future studies could use the 34 mg threshold when comparing ENMO derived ST estimates group level estimates previously published using the 100 count·min⁻¹ approach though it is important to consider the observed sensitivity and specificity results when interpreting findings.

Practical Implications

- Many previous studies estimated children's sedentary time using waist-mounted ActiGraph accelerometers and the 100 count·min⁻¹ threshold.
- The backward comparability of wrist-worn raw acceleration derived sedentary time estimates with the wealth of data collected using waist-mounted ActiGraphs is unknown.
- This study found that the 34 mg threshold could be applied to wrist accelerometer data to provide estimates of sedentary time that are equivalent to the ActiGraph waist-worn 100 count·min⁻¹ on average at the group level, though the sensitivity and specificity values observed in this study should be considered when interpreting findings.

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403 Table and Figure Legends

404 Table 1. Mean (SD) anthropometric, wear time and number of days included within analysis for boys
 405 and girls

	Boys N = 43		Girls N = 65	
	Mean or	SD	Mean or	SD
	Frequency		Frequency	
Age (y)	10	0.4	10	0.3
Height (cm)	139.5	7.9	138	7.4
Body mass (kg)	35.6	8.2	34.2	8.6
BMI (kg·m ²)	18.2	3.00	17.8	3.2
ActiGraph weekday wear (min·day ⁻¹)	739.9	115.6	738.8	100.4
ActiGraph weekend day wear (min·day ⁻¹)	631.8	110.8	661.5	108.3
ActiGraph valid weekdays included	41	N/A	64	N/A
ActiGraph valid weekend days included	30	N/A	46	N/A
Total valid included days	71	N/A	110	N/A

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410 Table 2. Comparisons between the ActiGraph <100 count·min⁻¹ standard and GA thresholds

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Criterion	GA threshold	Sedentary time (mins/included days)	Mean Bias (mins)	Bias vs Mean ST correlation	Cronbach's Alpha	ICC (Single measures)	Limits of agreement Lower	Limits of Agreement upper	Mean absolute percent error % (SD)	% agreement	Kappa
ActiGraph ≤100 count·min ⁻¹		492.4									
	23mg	342.6	149.8	-.11	0.88	0.59	-88.38	387.94	35.5 (18.2)	68	0.31
	30mg	440.4	52	-.33**	0.91	0.82	-166.41	270.45	20.2 (16.5)	69.4	0.36
	36mg	510.2	-17.8	-.49**	0.93	0.86	-230.47	194.81	15.8 (15.7)	69.8	0.38
	40mg	554.8	-62.4	-.58**	0.93	0.83	-276.22	151.36	18.2 (16.3)	69.9	0.39
	45mg	603.6	-111.2	-.67**	0.93	0.77	-327.29	104.95	25.1 (16.4)	69.5	0.39
	52mg	660	-167.6	-.75**	0.93	0.69	-391.09	55.91	34.9 (16.7)	68.8	0.39
	56mg	692.3	-199.9	-.79**	0.93	0.63	-430.48	30.7	40.7 (18)	68.2	0.38

412

414 Table 3. Comparisons between the ActiGraph ≤100 count·min⁻¹ standard and 34 mg and 35 mg GA thresholds.

415

Criterion	GA threshold	Sedentary time (mins)	Mean Bias (mins)	Bias vs Mean ST correlation	Cronbach's Alpha	ICC (Single measures)	Limits of agreement Lower	Limits of Agreement upper	Mean absolute percent error % (SD)	% agreement	Kappa	Sensitivity %	Specificity %
ActiGraph ≤100 count·min ⁻¹		492.4											
	34 mg	489.6	2.8	-.38**	0.92	0.86	-219.51	216.18	16.2 (15.8)	69.8	0.38	63.6	74.2
	35 mg	501	-8.6	-.47**	0.93	0.86	-220.88	203.71	15.8 (15.7)	69.8	0.38	64.8	73.4
Boys													
ActiGraph ≤100 count·min ⁻¹		499.7											
	34 mg	494.7	5	-.47**	0.93	0.87	-236.3	246.3	19 (16.8)	70.4	0.40	65.6	74.2
	35 mg	505.5	-5.8	-.52**	0.93	0.87	-247.3	235.6	19 (16.2)	70.4	0.40	66.7	73.3
Girls													
ActiGraph ≤100 count·min ⁻¹		487.6											
	34 mg	486.1	1.4	-.31**	0.92	0.85	-193.2	196	14.4 (14.9)	69.4	0.37	62.3	74.2
	35 mg	498	-10.4	-.42**	0.92	0.86	-202.9	182.1	13.7 (15)	69.4	0.37	63.6	73.5

416 Figure 1. ActiGraph ≤ 100 count \cdot min $^{-1} \pm 5\%$ zone of equivalence (467.7 minutes- 517 minutes, dotted
417 lines) and 90% confidence intervals for the GENEActiv sedentary time estimates classified using nine
418 thresholds

